# Outside Options in the Labor Market 

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## Motivation

In standard models of the labor market workers' wages depend on (typically unobserved) outside options

- Perfect competition: equally attractive option always exists $\Longrightarrow w=M P$
- Reality: next best option could vary in location, skill requirements, etc.

Outside job opportunities could vary across workers

- Could generate lower wages even for equally productive workers
- Ex: Women may have fewer options on average if they are less willing or able to commute

Challenge: Outside options are typically unobserved

## This Paper

Develop a method to estimate workers' outside employment opportunities

- Adapt standard marriage market models for use in the labor market (Becker 1973, Shapley-Shubik 1971)
- From this model, derive a sufficient statistic for outside options: Outside Options Index (OOI)
- "Concentration" index: learn about outside options from equilibrium outcomes of similar workers

Apply this model to German linked employer-employee data

1. Estimate empirical link between OOI and wage using a standard shift-share instrument

- $10 \%$ more options $\Longrightarrow 1.7 \%$ higher wages

2. $20 \%$ of gender gap is driven by differences in OOI (all coming from distance)

## Related Literature

## 1. Matching Models With Transfers

- Shapely \& Shubik (1971), Becker (1973), Ekeland, Heckman \& Nesheim (2004), Choo \& Siow (2006), Dupuy \& Galichon (2014)

2. Labor Market Imperfections and Wage Gaps

- Robinson (1933), Black (1995), Manning (2003) , Ransom \& Oaxaca (2010), Hirsch et al. (2010), Beaudry, Green \& Sand (2012), Hsieh et al. (2013), Bidner \& Sand (2016), Card, Cardoso \& Kline (2016), Card, Cardoso, Heining \& Kline (2018), Lamadon, Mogstad \& Setzler (2019)


## 3. Definition of a Labor Market

- Manning \& Petrongolo (2017), Nimczik (2018)


## 4. Labor Market Concentration

- Handwerker \& Spletzer (2015), Marinescu et al. (2018), Benmelech et al. (2018), Berger et al. (2019), Jarosch, Nimczik \& Sorkin (2019), Berger, Herkenhoff \& Mongey (2020), Schubert, Stansbury \& Taska (2020)

Theory

Empirical Setting and Data

Heterogeneity in Outside Options

Outside Options and Wage Inequality

## Matching Model with Two-Sided Heterogeneity

Continuum of workers of mass $\mathscr{I}=1$ and one-job firms of mass $\mathscr{J}=1$
If matched to firm $j$, worker $i$ produces

$$
\underbrace{\tau_{i j}}_{\text {total value }}=\underbrace{y_{i j}}_{\text {output }}+\underbrace{a_{i j}}_{\text {amenities }}
$$

Wages are used to transfer utility


## Equilibrium

Solve as a cooperative game (Shapley Shubik 1971)

- Static framework
- Perfect information

Equilibrium consists of an allocation $M$ and transfer $w_{i j}$ for each $(i, j) \in M$ which satisfies Details

$$
\begin{equation*}
\forall i^{\prime} \in \mathcal{I}, j^{\prime} \in \mathcal{J} \quad: \quad \omega_{i^{\prime}, m\left(i^{\prime}\right)}+\pi_{m^{-1}\left(j^{\prime}\right), j^{\prime}} \geq \tau_{i^{\prime} j^{\prime}} \tag{1}
\end{equation*}
$$

- Workers must earn more than they could elsewhere
- Firms must earn more than they could by hiring a different worker
- Compensation depends on distributions of productivity (y) and preferences (a)


## Functional Form Assumptions

1. Workers and jobs can be characterized by characteristics $\mathcal{X} \subseteq \mathbb{R}^{d_{x}}$ and $\mathcal{Z} \subseteq \mathbb{R}^{d_{z}}$

- Notation: worker $i$ has characteristics $X_{i}$ (density: $d\left(X_{i}\right)$ ) \& firm $j$ has characteristics $Z_{j}$ (density: $g\left(Z_{j}\right)$ )


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2. Allow for idiosyncratic preferences (Choo \& Siow, 2006, Dupuy \& Galichon, 2014)

$$
\tau_{i j}=\tau\left(x_{i}, z_{j}\right)+\epsilon_{i, z_{j}}+\varepsilon_{j, x_{i}}
$$

$2.1 \varepsilon \sim$ come from continuous logit models with scale $\alpha_{x}, \alpha_{z}$ Details

- Allows us to account for continuous observed characteristics (e.g. distance)
- Similar to standard MNL logit - but $\omega \nrightarrow \infty$ as (Cosslett 1988; Dagsvik 1994)
$2.2 \varepsilon_{i, z_{j}} \perp \varepsilon_{j, x_{i}}$
- Rules out interactions between worker/firm preferences


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- Rules out interactions between worker/firm preferences

IIA: Unobserved taste for jobs in an neighborhood of $z$ uncorrelated with unobserved taste for jobs in a neighborhood of $z^{\prime} \neq z$

## Outside Options and Compensation

In equilibrium [Proofs in Appendix A.5]:

1. Workers (employers) get "their" $\varepsilon_{i, z_{j}}\left(\varepsilon_{j, x_{i}}\right)$

$$
\omega_{i j}=\omega\left(x_{i}, z_{j}\right)+\epsilon_{i, z_{j}}, \quad \pi_{i j}=\pi\left(x_{i}, z_{j}\right)+\varepsilon_{j, x_{i}}
$$

2. The systematic portion of workers' compensation satisfies

$$
\omega(x, z)=\frac{\alpha_{x}}{\alpha_{x}+\alpha_{z}}(\underbrace{E\left[\omega \mid x_{i}\right]}_{\text {Expected Compensation }})+\frac{\alpha_{z}}{\alpha_{x}+\alpha_{z}}(\underbrace{\tau(x, z)-E[\pi \mid z]}_{\text {firm"rents" }})
$$

Note: $\frac{\alpha_{z}}{\alpha_{x}+\alpha_{z}}$ is larger when workers' idiosyncratic preferences are more variable than firms'

## Outside Options and Compensation

We can also decompose worker i's expected equilibrium compensation:


- Assuming firm profits stay constant, the OOI is a sufficient statistic for the effect of outside options on wages [Appendix A.3]


## Definition of Outside Options Index (OOI)

OOI is $E\left[\varepsilon_{i, z^{*}} \mid x_{i}\right]$ de-scaled

$$
\text { OOI }_{i}=\alpha_{z}^{-1} E\left[\varepsilon_{i, z^{*}} \mid x_{i}\right]=-\int f_{Z \mid X}\left(z_{j} \mid x_{i}\right) \log \frac{f_{Z \mid X}\left(z_{j} \mid x_{i}\right)}{g\left(z_{j}\right)}
$$

- Expected equilibrium value of $\epsilon_{i, z_{j}}$ for workers with characteristics $x_{i}$
- Concentration index that depends on both discrete and continuous characteristics
- Varies across workers due to differences in both preferences and skill (captured in $x_{i}$ )
- May vary across workers with identical $x_{i}$ due to labor market conditions (available $z_{j}$ )
- Nests transition-based measures (use discrete $X_{i}, Z_{j}$ based on industry/occupation)


## An Aside on Size-Based Market Power

Recent interest in the role of size-based monopsony power in determining wage mark-downs

In the paper [Appendix A.5] we present an extended model that allows for

- endogenous entry
- firms with multiple jobs

Key results:

- One-job case remains the upper bound for wages; a lower bound is set by assuming firms do not compete with themselves
- The expected difference in these bounds depends on how jobs are distributed across firms

$$
E\left[\overline{\omega_{i j}}-\underline{\omega_{i j}}\right]=-\sum_{k} \log \left(1-p_{k, i}\right)
$$

## Estimation: Assumptions

$$
O O I_{i}=-\int_{j} f_{j}^{i} \log f_{j}^{i}
$$

- where $f_{j}^{i}$ is the probability that $i$ works in job $j$.


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$$

- where $f_{j}^{i}$ is the probability that $i$ works in job $j$.

Assumption: Parameterization (Dupuy \& Galichon, 2014)

$$
\log \frac{f_{Z \mid X}\left(z_{j} \mid x_{i}\right)}{g\left(z_{j}\right)}=x_{i} A z_{j}+a\left(x_{i}\right)+b\left(z_{j}\right)
$$

where $a\left(X_{i}\right), b\left(Z_{j}\right)$ fix the marginal distributions
OOI is an index of concentration

- Estimated using cross-sectional distribution of similar workers
- On all observable dimensions
- Common index for unpredictability


## Estimating OOI

1. Simulate observations from $f\left(X_{i}\right) f\left(Z_{j}\right)$ and define

$$
Y= \begin{cases}1 & \text { Real Match } \\ 0 & \text { Simulated Match }\end{cases}
$$

2. Estimate a Logit model to recover $f_{j}^{i}$

$$
\begin{aligned}
\log \frac{P(Y=1 \mid X=x, Z=z)}{P(Y=0 \mid X=x, Z=z)} & =x A z+a(x)+b(z) \\
& =\frac{f\left(x_{i}, z_{j} \mid Y=1\right)}{f\left(x_{i}, z_{j} \mid Y=0\right)} \frac{P(Y=0)}{P(Y=1)} \\
& =\frac{f\left(x_{i}, z_{j}\right)}{f\left(x_{i}\right) f\left(z_{i}\right)}=f_{j}^{i} \cdot c
\end{aligned}
$$

3. Calculate $\widehat{f_{j}^{i}}$ for every possible worker-job combination and plug in

$$
\widehat{O O I_{i}}=\sum_{j} \widehat{f_{j}^{i}} \log \widehat{f_{j}^{i}}
$$

Theory

## Empirical Setting and Data

Heterogeneity in Outside Options

Outside Options and Wage Inequality

## Application: Germany

LIAB Longitudinal

- ~ $1 \%$ German workforce
- Cross-section: employed on 06/30/2014
- Focus on workers between 25 \& 55
- Supplement with task data from BIBB ( ${ }^{\sim}$ German $\mathrm{O}^{*}$ Net)
- Exploit linked establishment surveys


## Descriptive Statistics

|  | All |  | Male |  | Female |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean <br> (1) | SD <br> (2) | Mean (3) | SD <br> (4) | Mean <br> (5) | SD <br> (6) |
| Workers |  |  |  |  |  |  |
| Age | 46.32 | (11.64) | 45.89 | 11.87 | 46.82 | 11.34 |
| Female | 46\% | (0.50) | 0\% | --- | 1.00 | --- |
| German Citizen | 98\% | (0.14) | 98\% | 0.16 | 0.99 | (0.12) |
| Higher Secondary Degree | 28\% | (0.20) | 27\% | (0.20) | 29\% | (0.20) |
| Intermediate Secondary Degree | 31\% | (0.21) | 27\% | (0.20) | 34\% | (0.23) |
| Lower Secondary Degree | 19\% | (0.16) | 19\% | (0.15) | 21\% | (0.16) |
| Intermediate/Lower Education | 22\% | (0.17) | 27\% | (0.20) | 16\% | (0.14) |
| Daily Earnings | 87.30 | (51.23) | 104.27 | (50.87) | 67.3 | (43.90) |
| Distance | 12.90 | (39.15) | 15.80 | (43.71) | 9.49 | (32.64) |
| Jobs |  |  |  |  |  |  |
| Establishment size | 1547.75 | (7665.13) | 2183.74 | (9368.63) | 797.77 | (4847.42) |
| Sales per worker in 2013 (€) | 165341 | (187464.80) | 193785 | (199633.30) | 131798 | (165859.70) |
| Part-time contract | 31\% | (0.46) | 12\% | (0.33) | 53\% | (0.50) |
| Observations | 411,408 |  | 262,995 |  | 148,413 |  |

## Women Work Closer To Home

|  | Distance <br> (Miles) <br> $(1)$ | $<5$ Miles <br> $(2)$ | $5-20$ <br> Miles <br> $(3)$ | $20-50$ <br> Miles <br> $(4)$ | $50+$ <br> Miles <br> $(5)$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| All | 12.9 | $73.45 \%$ | $15.51 \%$ | $6.34 \%$ | $4.71 \%$ |
|  |  |  |  |  |  |
| Male | 15.8 | $69.28 \%$ | $17.23 \%$ | $7.37 \%$ | $6.11 \%$ |
| Female | 9.5 | $78.36 \%$ | $13.48 \%$ | $5.13 \%$ | $3.02 \%$ |
|  |  |  |  |  |  |
| Higher Secondary Degree | 22.1 | $62.50 \%$ | $19.42 \%$ | $9.10 \%$ | $8.98 \%$ |
| Intermediate Secondary Degree | 9.9 | $77.05 \%$ | $13.97 \%$ | $5.76 \%$ | $3.20 \%$ |
| Lower Secondary Degree | 9.4 | $77.78 \%$ | $13.46 \%$ | $5.58 \%$ | $3.18 \%$ |
| Intermediate/Lower Education | 8.0 | $79.04 \%$ | $14.42 \%$ | $4.08 \%$ | $2.48 \%$ |

## Baseline Measure of OOI

- $X_{i}$ : quadratic in age, female, PCA components for training occupation PCA
- $Z_{j}$ :
- Indicators for part-time/full-time, temp agency job, fixed term contract
- PCA components for occ \& industry, indicators for occupational complexity
- Establishment characteristics: size, share of females in management
- PCA based on establishment survey: business performance, investments, working hours, firm training, vocational training, "general"
- Distance: miles between worker's previous residence to establishment (400 districts)

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Theory
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Empirical Setting and Data

Heterogeneity in Outside Options

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## Distribution of OO



## Mass Layoff Exercise

- Involuntary job separations force workers to move to their outside options
- We use mass layoffs to show that the OOI is a meaningful measure of outside options
- We focus on workers who:
- Separated from their establishment between 1993-2014
- At an establishments with at least 50 workers
- At an establishments whose workforce declined $30 \%$ over the year
- With at least 3 years of tenure pre mass-layoff
- Are below age 55


## Mass Layoff Sample

|  | Main Sample |  | Mass Layoff Sample |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Mean <br> （1） | $\begin{aligned} & \hline \text { SD } \\ & (2) \\ & \hline \end{aligned}$ | Mean <br> （3） | $\begin{aligned} & \text { SD } \\ & (4) \\ & \hline \end{aligned}$ |
| Workers |  |  |  |  |
| Age | 46.32 | （11．64） | 38.64 | （10．62） |
| Female | 0.46 | （0．50） | 0.40 | （0．49） |
| German Citizen | 0.98 | （0．14） | 0.98 | （0．14） |
| Higher Secondary Degree | 28\％ | （0．20） | 18\％ | （0．15） |
| Intermediate Secondary Degree | 31\％ | （0．21） | 23\％ | （0．18） |
| Lower Secondary Degree | 19\％ | （0．16） | 20\％ | （0．16） |
| Intermediate／Lower Education | 22\％ | （0．17） | 39\％ | （0．24） |
| Daily Earnings | 87.30 | （51．23） | 66.35 | （85．93） |
| Workers | 411，408 |  | 13，404 |  |

## Outside Options and Mass Layoffs

- We compare workers within the same mass-layoff event $\psi_{j(i), t}$
- With different $\mathrm{OOI}_{i}$

$$
\begin{align*}
\widetilde{w}_{i, t}=\frac{w_{i, t}}{w_{i, 0}} & =\sum_{\tau=0}^{36} \lambda_{\tau} O O I_{i}+\psi_{j(i), t}+\mu_{t} X_{i t}+\nu_{i, t}  \tag{2}\\
e_{i, t} & =\sum_{\tau=0}^{36} \lambda_{\tau}^{\mathrm{emp}} O O I_{i}+\psi_{j(i), t}^{\mathrm{emp}}+\mu_{t}^{e m p} X_{i t}+\nu_{i, t}^{\mathrm{emp}} \tag{3}
\end{align*}
$$

## Mass Layoffs and Relative Wages

|  | (1) |  | (2) |  | (3) |  | (4) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 3 Months ( $\lambda_{3}$ ) | 0.071 | *** | 0.071 | *** | 0.067 | *** | 0.068 | *** |
|  | (0.022) |  | (0.022) |  | (0.023) |  | (0.023) |  |
| 6 Months ( $\lambda_{6}$ ) | 0.089 | *** | 0.089 | *** | 0.083 | *** | 0.083 | *** |
|  | (0.024) |  | (0.024) |  | (0.026) |  | (0.027) |  |
| 12 Months ( $\lambda_{12}$ ) | 0.103 | *** | 0.102 | *** | 0.089 | *** | 0.088 | *** |
|  | (0.027) |  | (0.027) |  | (0.031) |  | (0.031) |  |
| 24 Months ( $\lambda_{24}$ ) | 0.109 | *** | 0.109 | *** | 0.079 | ** | 0.075 | ** |
|  | (0.034) |  | (0.034) |  | (0.036) |  | (0.036) |  |
| Establishment-Month FE | $\checkmark$ |  | $\checkmark$ |  | $\checkmark$ |  | $\checkmark$ |  |
| Tenure |  |  | $\checkmark$ |  | $\checkmark$ |  | $\checkmark$ |  |
| Age |  |  |  |  | $\checkmark$ |  | $\checkmark$ |  |
| Education |  |  |  |  | $\checkmark$ |  | $\checkmark$ |  |
| Gender |  |  |  |  | $\checkmark$ |  | $\checkmark$ |  |
| Training Occupation Characteristics |  |  |  |  |  |  | $\checkmark$ |  |
| Observations | 547,353 |  | 547,353 |  | 547,353 |  | 547,353 |  |
| Workers | 13,404 |  | 13,404 |  | 13,404 |  | 13,404 |  |

## Mass Layoffs and Employment

|  | (1) |  | (2) |  | (3) | (4) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 3 Months ( $\lambda_{3}$ ) | $\begin{gathered} \hline 0.016 \\ -(0.005) \end{gathered}$ | *** | $\begin{gathered} \hline 0.016 \\ -(0.005) \end{gathered}$ | *** | $\begin{array}{cc} \hline 0.013 & * * \\ -(0.006) & \end{array}$ | $\begin{array}{cc} \hline 0.012 \\ -(0.006) \end{array} \quad \text { ** }$ |
| 6 Months ( $\lambda_{6}$ ) | $\begin{gathered} 0.008 \\ -(0.006) \end{gathered}$ |  | $\begin{gathered} 0.008 \\ -(0.006) \end{gathered}$ |  | $\begin{gathered} 0.004 \\ -(0.006) \end{gathered}$ | $\begin{gathered} 0.002 \\ -(0.006) \end{gathered}$ |
| 12 Months ( $\lambda_{12}$ ) | $\begin{gathered} 0.016 \\ -(0.006) \end{gathered}$ | ** | $\begin{gathered} 0.016 \\ -(0.006) \end{gathered}$ | ** | $\begin{gathered} 0.009 \\ -(0.007) \end{gathered}$ | $\begin{gathered} 0.007 \\ -(0.007) \end{gathered}$ |
| 24 Months ( $\lambda_{24}$ ) | $\begin{gathered} 0.017 \\ -(0.007) \end{gathered}$ |  | $\begin{gathered} 0.017 \\ -(0.007) \end{gathered}$ |  | $\begin{gathered} 0.011 \\ -(0.007) \end{gathered}$ | $\begin{gathered} 0.007 \\ -(0.007) \end{gathered}$ |
| Establishment-Month FE | $\checkmark$ |  | $\checkmark$ |  | $\checkmark$ | $\checkmark$ |
| Tenure |  |  | $\checkmark$ |  | $\checkmark$ | $\checkmark$ |
| Age |  |  |  |  | $\checkmark$ | $\checkmark$ |
| Education |  |  |  |  | $\checkmark$ | $\checkmark$ |
| Gender |  |  |  |  | $\checkmark$ | $\checkmark$ |
| Training Occupation Characteristics |  |  |  |  |  | $\checkmark$ |
| Observations | 547,353 |  | 547,353 |  | 547,353 | 547,353 |
| Workers | 13,404 |  | 13,404 |  | 13,404 | 13,404 |

## Geographic Variation



## Distribution of the OO

|  |  |  | Quantiles |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | Mean | SD | 25th | 50 th | 75 th |
|  | $(1)$ | $(2)$ | $(3)$ | $(4)$ | $(5)$ |
| All | -4.82 | 0.97 | -5.37 | -4.70 | -4.14 |
|  |  |  |  |  |  |
| Male | -4.74 | 1.00 | -5.28 | -4.59 | -4.05 |
| Female | -4.92 | 0.91 | -5.47 | -4.83 | -4.27 |
|  |  |  |  |  |  |
| Citizen | -4.82 | 0.95 | -5.36 | -4.70 | -4.14 |
| Non-Citizen | -5.10 | 1.37 | -5.52 | -4.86 | -4.34 |
|  |  |  |  |  |  |
| Higher Secondary Degree | -4.58 | 0.92 | -5.01 | -4.45 | -3.99 |
| Intermediate Secondary Degree | -4.76 | 0.87 | -5.32 | -4.67 | -4.11 |
| Lower Secondary Degree | -4.91 | 0.95 | -5.47 | -4.80 | -4.22 |
| Intermediate/Lower Education | -5.14 | 0.93 | -5.69 | -5.08 | -4.46 |

## Heterogeneity in the OOI

|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Female | $\begin{aligned} & \mathbf{- 0 . 2 9 5}^{* * *} \\ & (0.009) \end{aligned}$ | $\begin{aligned} & \mathbf{- 0 . 2 6 8}^{* * *} \\ & (0.011) \end{aligned}$ | $\begin{aligned} & \mathbf{- 0 . 2 8 3}^{* * *} \\ & (0.007) \end{aligned}$ | $\begin{aligned} & \mathbf{- 0 . 2 5 5}^{* * *} \\ & (0.008) \end{aligned}$ | $\begin{aligned} & \hline-0.201^{* * *} \\ & (0.008) \end{aligned}$ | $\begin{aligned} & \hline-0.237^{* * *} \\ & (0.008) \end{aligned}$ | $\begin{aligned} & \hline-0.344 \\ & (0.009) \end{aligned}$ |
| Non-Citizen | $\begin{aligned} & -0.262^{* * *} \\ & (0.036) \end{aligned}$ | $\begin{aligned} & -0.226^{* * *} \\ & (0.032) \end{aligned}$ | $\begin{aligned} & -0.5533^{* * *} \\ & (0.030) \end{aligned}$ | $\begin{aligned} & -0.498^{* * *} \\ & (0.026) \end{aligned}$ | $\begin{aligned} & -0.539^{* * *} \\ & (0.022) \end{aligned}$ | $\begin{aligned} & -0.494^{* * *} \\ & (0.020) \end{aligned}$ | $\begin{aligned} & -0.675 \quad * * * \\ & (0.025) \end{aligned}$ |
| Lower-Secondary Certificate | $\begin{aligned} & -0.601 \quad * * * \\ & (0.014) \end{aligned}$ | $\begin{aligned} & -0.535^{* * *} \\ & (0.014) \end{aligned}$ | $\begin{aligned} & -0.526^{* * *} \\ & (0.011) \end{aligned}$ | $\begin{aligned} & -0.474^{* * *} \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.504^{* * *} \\ & (0.011) \end{aligned}$ | $\begin{aligned} & -0.464^{* * *} \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.374 \quad * * * \\ & (0.010) \end{aligned}$ |
| Intermediate | $\begin{aligned} & -0.2366^{* * *} \\ & (0.011) \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.211^{* * *} \\ & (0.011) \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.110^{* * *} \\ & (0.008) \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.110^{* * *} \\ & (0.008) \end{aligned}$ | $\begin{aligned} & -0.129^{* * *} \\ & (0.009) \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.129^{* * *} \\ & (0.008) \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.098 \quad * * * \\ & (0.009) \\ & \hline \end{aligned}$ |
| Age Controls | Quadratic | Quadratic | Quadratic | Quadratic | Quadratic | Quadratic | Quadratic |
| Training Occupation FE |  | $\checkmark$ |  | $\checkmark$ |  | $\checkmark$ | $\checkmark$ |
| District FE |  |  | $\checkmark$ | $\checkmark$ |  |  | $\checkmark$ |
| Establishment FE |  |  |  |  | $\checkmark$ | $\checkmark$ |  |
| OOI Based on Vacancies |  |  |  |  |  |  | $\checkmark$ |
| Adjusted R-Squared | 0.133 | 0.253 | 0.530 | 0.629 | 0.573 | 0.627 | 0.562 |
| Observations | 375,765 | 375,765 | 375,765 | 375,765 | 375,765 | 375,765 | 375,765 |

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Theory
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Empirical Setting and Data

Heterogeneity in Outside Options

Outside Options and Wage Inequality

## Linking OOI and Wages

$$
\log w_{i}=\alpha O O I_{i}+\beta X_{i}+\varepsilon_{i}
$$

1. Endogeneity: OOI is an equilibrium object, correlated with worker productivity
2. Measurement error: OOI is measured with noise

Measure link between outside options and wages using instruments that change workers' option sets

- Ideal instrument holds firm profits constant
- Use a standard shift-share instrument, explore robustness with exporting firms


## Shift-Share OOI

Idea: Compare workers in the same industry with outside options in different industries
Specification: Look at change in wages 2004-2014 within industries ( $j$ )
Instrument Construction

$$
\begin{array}{ccc}
\Delta_{04}^{14} \log w_{i} & = & \alpha \Delta_{04}^{14} O O I_{i} \\
\Delta_{04}^{14} \text { OOI }_{i} & =\gamma Z_{j(i, 2004), r(i, 2004)} & +\beta \Delta_{04}^{14} X_{i}+\operatorname{Ind}_{j(i, 2004)}^{14} X_{i}+v_{i} \operatorname{Ind}_{j(i, 2004)}+\epsilon_{i} \tag{4}
\end{array}
$$

where $Z_{j}$ is the expected change in OOI for individuals in industry $j$ and region $r$ in 2004
ID: exogeneity of shocks

$$
E\left[\varepsilon_{i} Z_{j(i, 2004), r(i, 2004)} \mid \operatorname{Ind} d_{j(i)}^{04}, \Delta_{04}^{14} X_{i}\right]=0
$$

## Shift－Share OOI：Instrument Details

1．Calculate the predicted OOI for each individual

$$
\left.\widetilde{O O I_{i, 2014}}=-\sum_{z_{j}} f_{Z \mid X} \widehat{\left(z_{j} \mid x_{i}\right.}\right)\left(\frac{\log \widehat{f_{Z \mid X}\left(z_{j} \mid x_{i}\right)}}{\log \widetilde{g}_{14}\left(z_{j}\right)}\right)
$$

2．Calculate the predicted change in OOI

$$
\Delta_{04}^{14} \widetilde{O O I_{i}}=\widetilde{O O I_{i, 2014}}-O O I_{i, 2004}
$$

3．Average across individuals in region $j$ and industry $r$ in 2004

$$
Z_{j, r}=\frac{1}{|\mathcal{S}(j, r)|} \sum_{i \in \mathcal{S}(j, r)} \Delta_{04}^{14} \widetilde{O_{i}}
$$

## Shift-Share Results

|  | Full Sample |  |  | By Exporting Share of Sales |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | $\begin{gathered} \text { More than } \\ 33 \% \\ \hline \end{gathered}$ | Between 1 and $33 \%$ | $\begin{gathered} \hline \text { Less than } \\ 1 \% \\ \hline \end{gathered}$ |
|  | (1) | (2) | (3) | (4) | (5) | (6) |
| First Stage | $\begin{array}{cc} \hline 0.299^{* * *} \\ (0.064) & \end{array}$ | $\begin{array}{cl} 0_{0.276} & * * * \\ (0.048) & \end{array}$ | $\begin{array}{cc} 0_{0.242} & \text { *** } \\ (0.064) & \end{array}$ | $\begin{array}{cc} 0_{0.353} & \text { *** } \\ (0.104) & \end{array}$ | $\begin{array}{cc} \hline 0.204 & * * * \\ (0.059) & \end{array}$ | $\begin{array}{cc} \hline 0.272 & * * * \\ (0.080) & \end{array}$ |
| Reduced Form | $\begin{aligned} & 0.0517 \text { ** } \\ & (0.021) \end{aligned}$ | $\begin{aligned} & 0.0504^{* *} \\ & (0.021) \end{aligned}$ | $\begin{gathered} 0.038 \\ (0.024) \end{gathered}$ | $\begin{gathered} 0.080 \\ (0.026) \end{gathered}$ | $\begin{gathered} 0.009 \\ (0.026) \end{gathered}$ | $\begin{gathered} 0.031 \\ (0.023) \end{gathered}$ |
| 2SLS | $\begin{aligned} & 0.173^{* * *} \\ & (0.063) \\ & \hline \end{aligned}$ | $\begin{aligned} & 0.183 \quad * * * \\ & (0.068) \\ & \hline \end{aligned}$ | $\begin{gathered} 0.156 \\ (0.092) \\ \hline \end{gathered}$ | $\begin{array}{ll} 0.227 & \text { *** } \\ (0.071) & \end{array}$ | $\begin{gathered} 0.046 \\ (0.123) \\ \hline \end{gathered}$ | $\begin{gathered} 0.114 \\ (0.096) \\ \hline \end{gathered}$ |
| Industry FE | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ |
| Age Controls | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ |
| Demographic Controls |  | $\checkmark$ | $\checkmark$ |  |  |  |
| Regional Controls |  |  | $\checkmark$ |  |  |  |
| F (First Stage) | 21.95 | 32.82 | 14.5 | 11.52 | 12.04 | 11.56 |
| Number of industry-regions | 5510 | 5510 | 5510 | 2195 | 2525 | 790 |
| Observations | 435,586 | 435,586 | 435,586 | 144,039 | 147,529 | 144,018 |

## Shift-Share Heterogeneity

|  | By Gender |  | By Education |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Male | Female | Higher <br> Secondary | Intermediate Secondary | Lower <br> Secondary |
|  | (1) | (2) | (3) | (4) | (5) |
| First Stage | $\begin{array}{cc} \hline 0.309 & * * * \\ (0.080) & \end{array}$ | $\begin{aligned} & \mathbf{c o s}^{0.266}{ }^{* * *} \\ & (0.050) \end{aligned}$ | $\begin{array}{cc} \mathbf{c}_{0.232} & \text { *** } \\ (0.079) & \end{array}$ | $\begin{array}{cc} \hline 0.203 & * * * \\ (0.053) & \end{array}$ | $\begin{gathered} \mathbf{c o s}^{0.321}{ }^{* * *} \\ (0.049) \end{gathered}$ |
| Reduced Form | $\begin{aligned} & 0.0673^{* * *} \\ & (0.021) \end{aligned}$ | $\begin{gathered} 0.019 \\ (0.022) \end{gathered}$ | $\begin{gathered} 0.031 \\ (0.022) \end{gathered}$ | $\begin{aligned} & 0.046 \\ & (0.022) \end{aligned}$ | $\begin{aligned} & 0.080 \\ & (0.026) \end{aligned}$ |
| 2SLS | $\begin{array}{ll} 0.218 & * * * \\ (0.059) \end{array}$ | $\begin{gathered} 0.071 \\ (0.086) \end{gathered}$ | $\begin{gathered} 0.134 \\ (0.099) \end{gathered}$ | $\underbrace{0.228}(0.103) ~ * * ~ ? ~$ | $\begin{aligned} & 0.247 \quad \text { *** } \\ & (0.078) \end{aligned}$ |
| Industry FE | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ |
| Age Controls | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ |
| F (first stage) | 14.77 | 27.97 | 8.56 | 14.89 | 43.45 |
| Observations | 283,550 | 152,036 | 96,148 | 148,136 | 91,793 |

## Decomposing Wage Gaps

1. Baseline: Raw wage gap

$$
\log w_{i}=\beta_{0} X_{i}+\epsilon_{i}
$$

- Mincer regression of log wages on demographic characteristics: indicators for each education group, a quadratic function of age, gender, citizenship status, part-time indicators

2. Wage gap explained by the OOI:

$$
\log w_{i}=\underbrace{\widehat{\alpha}}_{.17} \mathrm{OOI}_{i}+\beta_{1} X_{i}+\nu_{i}
$$

3. Wage gap explained by commuting costs:

$$
\log w_{i}=\underbrace{\hat{\alpha}}_{.17}\left(\mathrm{OOI}_{i}-\widetilde{O O I}_{i}\right)+\beta_{2} X_{i}+\epsilon_{i}
$$

Note: to account for top-coding, we estimate each equation using a Tobit model $\qquad$

## Wage Gaps and Distance

- Assign everyone the "commuting cost" or a 40 year old male citizen with highest level of education



## Discussion

- Developed a method to estimate workers' outside employment opportunities
- Adapted standard marriage market models for use in the labor market (Becker 1973, Shapley-Shubik 1971)
- Derived a sufficient statistic for outside options: Outside Options Index (OOI)
- Applied this approach to linked employer-employee data from Germany
- Males, German citizens, urban residents have more options
- $10 \%$ more options yields $1.7 \%$ higher income
- Differences in options tend to increase between-group wage inequality: $20 \%$ of gender gap


## Thank You

Appendix

## Solution: Equilibrium

Stable equilibrium (core allocation) includes:

1. Allocation of workers and jobs $m: \mathcal{I} \rightarrow \mathcal{J}$
2. Transfers $w_{i j}$

Which satisfies the following conditions:

1. No profitable deviations $\forall i \in \mathcal{I}, \forall j \in \mathcal{J}$ :

2. Participation constraint

$$
\begin{array}{ll}
\forall i \in I & : \quad \omega_{i, m(i)} \geq u_{i} \\
\forall j \in J & :
\end{array} \pi_{m^{-1}(j), j} \geq v_{j}
$$

where $u_{i}, v_{j}$ are the value of unemployment or vacancy Returm

## Continuous Logit Assumptions

$$
\begin{array}{cc}
\tau_{i j}=\tau\left(x_{i}, z_{j}\right)+\varepsilon_{i, z_{j}}+\varepsilon_{j, x_{i}} \\
\text { s.t. } & \varepsilon_{i, z_{j}} \perp \varepsilon_{j, x_{i}} \\
& \varepsilon_{i, z_{j}}, \varepsilon_{j, x_{i}} \sim C L(\alpha)
\end{array}
$$

- Each worker (job) knows about a random subset of the available jobs (workers)
- For each of these jobs (workers), the relevant party draws $\epsilon$ from a Poisson process on $\mathcal{Z} \times \mathbb{R}$ with intensity

$$
f(z) d z \times e^{-\varepsilon} d \varepsilon
$$

- The maximum value on each Borel measurable subset is $E V_{1}$ with scale $\alpha$


## Continuous Logit Choice

$Q_{z_{j} \mid x_{i}}$ is the measure of $x_{i}$ times their share that chooses $z_{j}$.

$$
Q_{z_{j} \mid x_{i}}=f\left(x_{i}\right) f\left(z_{j} \mid x_{i}\right)
$$

In continuous logit the share to choose $z_{j}$ is

$$
\frac{\exp \omega\left(x_{i}, z_{j}\right) f\left(z_{j}\right)}{\int_{z^{\prime}} \exp \omega\left(x_{i}, z^{\prime}\right) f\left(z^{\prime}\right) d z^{\prime}}=\frac{\exp \omega\left(x_{i}, z_{j}\right) f\left(z_{j}\right)}{\exp E\left[\omega_{i} \mid x_{i}\right]}
$$

Market clears when

$$
\begin{gathered}
Q_{z_{j} \mid x_{i}}=\frac{\exp \omega\left(x_{i}, z_{j}\right) f\left(z_{j}\right) f\left(x_{i}\right)}{\exp E\left[\omega_{i} \mid x_{i}\right]}=\frac{\exp \pi\left(x_{i}, z_{j}\right) f\left(z_{j}\right) f\left(x_{i}\right)}{\exp E\left[\pi_{j} \mid z_{j}\right]}=Q_{x_{i} \mid z_{j}} \\
\omega\left(x_{i}, z_{j}\right)-\pi\left(x_{i}, z_{j}\right)=E\left[\omega_{i} \mid x_{i}\right]-E\left[\pi_{j} \mid z_{j}\right]
\end{gathered}
$$

By definition

$$
\omega\left(x_{i}, z_{j}\right)+\pi\left(x_{i}, z_{j}\right)=\tau\left(x_{i}, z_{j}\right)
$$

And the sum gives the solution

## PCA Components for Occupations

|  | N | First Component | Second Component |
| :--- | :---: | :--- | :--- |
| Hours | 11021 | Work on Sundays and public holidays | Hours per week like to work |
| Type of Task | 15035 | Have responsibility for other people | Cleaning, waste, recycling |
| Requirements | 10904 | Face acute pressure and deadlines | Highly specific regulations |
| Physical | 20036 | Oil, dirt, grease, grime | Pathogens, bacteria |
| Mental | 17790 | Support from colleagues | Often missing information about work |

## PCA Components from Estab. Survey (Z)

|  | N | First Component | Second Component |
| :--- | :---: | :--- | :--- |
| Business Performance | 8824 | Member of chamber of industry | Profit |
| Investment \& Innovation | 8824 | IT investment | Total investment |
| Hours | 8824 | Vacation credit policy | Flexible hours |
| Vocational Training | 8824 | Offer apprenticeship | Ability to fill training |
| General | 8824 | Family managed | Staff representation |

## Proof

$$
\begin{gathered}
f_{j}^{i}=f(j \mid i)=f\left(j \mid X=x_{i}\right)= \\
=f\left(j \mid Z=z_{j}, X=x_{i}\right) f\left(Z=z_{j} \mid X=x_{i}\right)= \\
=f\left(j \mid Z=z_{j}\right) \frac{f\left(X=x_{i}, Z=z_{j}\right)}{f\left(X=x_{i}\right)}= \\
=\frac{|J|^{-1}}{f\left(Z=z_{j}\right)} \frac{f\left(X=x_{i}, Z=z_{j}\right)}{f\left(X=x_{i}\right)}
\end{gathered}
$$

## Mass-Layoffs

- Outcome variable: Daily wage divided by baseline $\frac{w_{t}}{w_{0}}$


Mass-Layoffs: Relative Income


Mass Layoffs - Job Search


